

On the Variance of Quickselect

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- 1 Introduction
- 2 General results
- 3 The variance of median-of-three
- 4 The variance for large samples

Problem: Given an array A of n items and a rank m ,
 $1 \leq m \leq n$, find the m th smallest element in A .

The algorithm should work in (expected) linear time $\Theta(n)$,
irrespective of m .

Hoare (1962) invents **quickselect**: pick some element p from the array, called the **pivot**, rearrange the contents of A so that all elements in A smaller than p are to its left, and all elements larger than p are to its right; if p is at position $j = m$ it is the sought element; if $j > m$ proceed recursively in $A[1..j - 1]$, otherwise in $A[j + 1..n]$.

```
Elem quickselect(vector<Elem>& A, int m) {  
    int l = 0; int u = A.size() - 1;  
    int k, p;  
    while (l <= u) {  
        p = select_pivot(A, l, u, m);  
        swap(A[p], A[l]);  
        partition(A, l, u, j);  
        if (m < j) u = j-1;  
        else if (m > j) l = j+1;  
        else return A[j];  
    }  
}
```

Knuth (1971) shows that

$$\mathbb{E}[C_{n,m}] = 2(n+3 + (n+1)H_n - (m+2)H_m - (n+3-m)H_{n+1-m}),$$

with $H_n = \sum_{1 \leq i \leq n} (1/i) = \log n + \mathcal{O}(1)$ the n th harmonic number.

- The **expectation characteristic function**:

$$f(\alpha) = \lim_{\substack{n \rightarrow \infty \\ m/n \rightarrow \alpha}} \frac{\mathbb{E}[C_{n,m}]}{n}$$

- The **second factorial moment characteristic function**:

$$g(\alpha) = \lim_{\substack{n \rightarrow \infty \\ m/n \rightarrow \alpha}} \frac{\mathbb{E}[C_{n,m}^2]}{n^2}$$

- For the **variance** we have

$$v(\alpha) = \lim_{\substack{n \rightarrow \infty \\ m/n \rightarrow \alpha}} \frac{\mathbb{V}[C_{n,m}]}{n^2} = g(\alpha) - f^2(\alpha)$$

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Example

- Standard quickselect:

$$f(\alpha) = m_0(\alpha) = 2 - 2(\alpha \ln \alpha + (1 - \alpha) \ln(1 - \alpha)) = 2 + 2 \cdot \mathcal{H}(\alpha)$$

- Median-of-three:

$$f(\alpha) = m_1(\alpha) = 2 + 3\alpha(1 - \alpha)$$

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- Standard quickselect:

$$m_0(0) = m_0(1) = 2$$

$$m_0(1/2) = 2 + 2 \ln 2 \approx 3.386$$

- Median-of-three:

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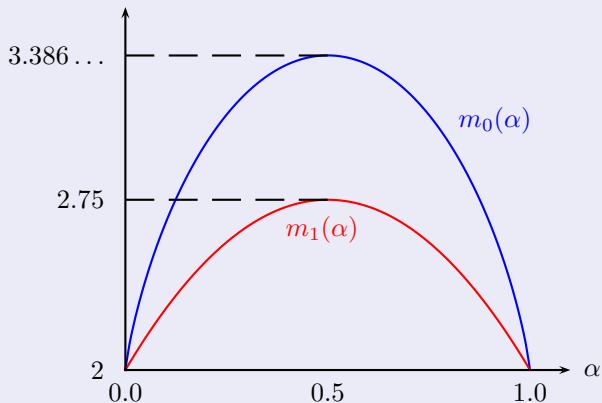
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A plot of the **standard quickselect** characteristic function versus **median-of-three** characteristic function



- **Adaptive sampling** uses a sample of s elements to choose a pivot for each recursive stage of quickselect.
- If the **current relative rank** is $\alpha = m/n$, we select the element of rank $r(\alpha)$ from the sample

Example

- Standard quickselect: $s = 1, r(\alpha) = 1$
- Median-of- $(2t + 1)$: $s = 2t + 1, r(\alpha) = t + 1$
- Proportion-from- s : $r(\alpha) \approx \alpha \cdot s$

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Example

We are looking the fourth element ($m = 4$) out of $n = 15$ elements

9	5	10	12	3	1	11	15	7	2	8	13	6	4	14
---	---	----	----	---	---	----	----	---	---	---	----	---	---	----

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We are looking the fourth element ($m = 4$) out of $n = 15$ elements

2	3	1	4	5	6	8	7	9	15	11	13	12	10	14
---	---	---	---	---	---	---	---	---	----	----	----	----	----	----

An adaptive sampling strategy can be characterized by the value of $r(\alpha)$ for a finite set of ℓ intervals that partition $[0, 1]$, i.e., $r_k = r(\alpha)$ if $\alpha \in I_k$, $1 \leq k \leq \ell$.

$$0 = a_0 < a_1 < a_2 < \cdots < a_{\ell-1} < a_\ell = 1,$$

$$I_1 = [0, a_1], \quad I_\ell = [a_{\ell-1}, 1],$$

$$I_k = (a_{k-1}, a_k] \quad \text{if } k > 1 \text{ and } a_k \leq 1/2,$$

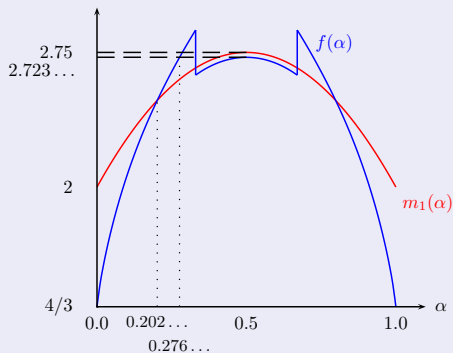
$$I_k = [a_{k-1}, a_k) \quad \text{if } k < \ell \text{ and } a_{k-1} > 1/2, \text{ and}$$

$$I_k = (a_{k-1}, a_k) \quad \text{if } a_{k-1} \leq 1/2 < a_k \text{ and } 1 < k < \ell.$$

Example

- Standard quickselect: $s = 1; \ell = 1; r_1 = 1$
- Median-of- $(2t + 1)$: $s = 2t + 1; \ell = 1; r_1 = t + 1$
- Proportion-from- s : $\ell = s; r_k = k$
- “Pure” proportion-from- s : proportion-from- $s + a_k = k/s$

A plot of **median-of-three** characteristic function versus **proportion-from-three** $f(\alpha)$



Theorem (Martínez, Panario, Viola (2004))

The expectation characteristic function $f(\alpha)$ of any adaptive sampling strategy satisfies

$$f(\alpha) = 1 + \frac{s!}{(r(\alpha) - 1)!(s - r(\alpha))!} \times \left[\int_{\alpha}^1 f(\alpha/x) x^{r(\alpha)} (1 - x)^{s - r(\alpha)} dx + \int_0^{\alpha} f\left(\frac{\alpha - x}{1 - x}\right) x^{r(\alpha) - 1} (1 - x)^{s + 1 - r(\alpha)} dx \right].$$

Lemma (Martínez, Panario, Viola (2004))

Let f_k be the restriction of $f(\alpha)$ to the k th interval I_k , and $r_k = r(\alpha)$ when $\alpha \in I_k$. For any adaptive sampling strategy

$$\begin{aligned} \frac{d^{s+2}}{d\alpha^{s+2}} f_k(\alpha) &= \frac{(-1)^{s+1-r_k} \cdot s!}{\alpha^{s+1-r_k} (r_k - 1)!} \frac{d^{r_k+1}}{d\alpha^{r_k+1}} f_k(\alpha) \\ &+ \frac{s!}{(1-\alpha)^{r_k} (s-r_k)!} \frac{d^{s+2-r_k}}{d\alpha^{s+2-r_k}} f_k(\alpha). \end{aligned}$$

Theorem (Martínez, Panario, Viola (2004))

*Proportion-from- s sampling with $s \rightarrow \infty$ achieves **optimal** expected performance:*

$$f(\alpha) = 1 + \min(\alpha, 1 - \alpha)$$

- 1 Introduction
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- 3 The variance of median-of-three
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Theorem

The second factorial moment characteristic function $g(\alpha)$ of any adaptive sampling strategy satisfies

$$g(\alpha) = 2f(\alpha) - 1 + \frac{s!}{(r(\alpha) - 1)!(s - r(\alpha))!} \left[\int_{\alpha}^1 g(\alpha/x) x^{r(\alpha)+1} (1-x)^{s-r(\alpha)} dx + \int_0^{\alpha} g\left(\frac{\alpha-x}{1-x}\right) x^{r(\alpha)-1} (1-x)^{s+2-r(\alpha)} dx \right].$$

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Lemma

For any adaptive sampling strategy

$$\lim_{\alpha \rightarrow 0} v(\alpha) = \frac{r_0(s+1)}{(s+1-r_0)((s+2)(s+1) - r_0(r_0+1))},$$

where $r_0 = \lim_{\alpha \rightarrow 0} r(\alpha)$.

Example

- Median-of- $(2t+1)$: $v(0) = v(1) = \frac{2}{3t+4}$
- Proportion-from- s : $v(0) = v(1) = \frac{s+1}{s^2(s+3)} \sim \frac{1}{s^2} + \mathcal{O}(s^{-3})$

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- 1 Introduction
- 2 General results
- 3 The variance of median-of-three**
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The differential equation to find the **expectation characteristic function** is

$$\frac{d^2\phi}{d\alpha^2} = 6 \left(\frac{1}{\alpha^2} + \frac{1}{(1-\alpha)^2} \right) \phi(\alpha)$$

with $\phi(\alpha) = f'''(\alpha)$

For the **second moment characteristic function** $g(\alpha)$ we have

$$\frac{d^2\phi}{d\alpha^2} = 6 \left(\frac{1}{\alpha^2} + \frac{1}{(1-\alpha)^2} \right) \phi(\alpha)$$

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That's exactly the same ODE as for $f(\alpha)$!!

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The independent term in the ODE for $g(\alpha)$ vanishes, since $f(\alpha) = 2 + 3\alpha(1 - \alpha)$ and $f^{(vi)}(\alpha) = 0$.

- We integrate four times the solution found
- We plug the general form back into the integral equation to determine the value of the arbitrary constants; we also use the symmetry of $g(\alpha)$
- The final solution is

$$\begin{aligned}
 g(\alpha) = & -\frac{288}{35}\alpha^2(\ln(\alpha) + \ln(1 - \alpha)) - \frac{288}{35}\ln(1 - \alpha) \\
 & + \frac{576}{35}\alpha\ln(1 - \alpha) + \frac{30}{7} - \frac{24}{245}\alpha^8 + \frac{96}{245}\alpha^7 \\
 & - \frac{48}{175}\alpha^6 - \frac{96}{175}\alpha^5 - \frac{48}{35}\alpha^4 + \frac{144}{35}\alpha^3 \\
 & - \frac{7332}{1225}\alpha^2 + \frac{132}{35}\alpha,
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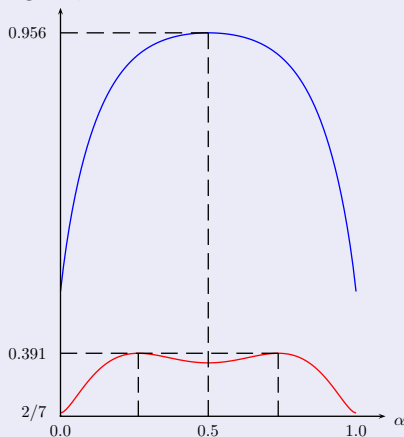
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A plot of $v(\alpha)$ for **standard quickselect** (Kirschenhofer, Prodinger (1998)) and for **median-of-three**



We've got the general form of $g(\alpha)$ for standard quickselect and proportion-from-2, but the process of determining the arbitrary constants is still not finished ...

It's much harder than we thought!!

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- Intuition: Using very large sample and proportion-from- s helps, because we get a very good pivot, very close to the sought element
- We should make sure that our pivot is very close **BUT** at the right side of the sought element! (i.e., slightly to the right if $\alpha < 1/2$, slightly to the left if $\alpha > 1/2$)

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Definition

A family of sampling strategies is *biased* if, for $\alpha < 1/2$,

$$r(\alpha) > s \cdot \alpha + 1 - \alpha$$

The proof of Martínez, Panario, Viola (2004) for adaptive optimal sampling works also for $s = s(n)$, as long as $s \rightarrow \infty$ and $s/n \rightarrow 0$ if $n \rightarrow \infty$.

$$\mathbb{E}[C_{n,m}] = n + \min(m, n - m) + \Theta\left(\max\left(s, \frac{n}{s}\right)\right)$$

Theorem

Biased proportion-from- s sampling with $s \rightarrow \infty$ has subquadratic variance:

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The same holds true for median-of- $(2t + 1)$, when $t \rightarrow \infty$

Theorem

For biased proportion-from- s sampling with increasing variable-sized samples (i.e., $s = s(n) \rightarrow \infty, s/n \rightarrow 0$), we have

$$\mathbb{V}[C_{n,m}] = \Theta \left(\max \left(\frac{n^2}{s}, n \cdot s \right) \right)$$

Theorem

The variance and the expected value of proportion-from- s , with variable-sized samples, is minimized when

$$s = \Theta(\sqrt{n})$$

Floyd and Rivest (1970) proposed an algorithm which uses sampling to obtain two pivots at each stage and achieves optimal expected performance.

However, the algorithm is more complicated and uses samples of size $\Theta(n^{2/3} \log n)$ (why!?)

Current work:

- Exact solutions for particular strategies (e.g., proportion-from-2)
- Precise asymptotic estimates of the optimal sample size when $s \rightarrow \infty$
- We need better estimates of the behavior when $s \rightarrow \infty$, e.g., we know that $f(\alpha) = 1 + \min(\alpha, 1 - \alpha) + \mathcal{O}(s^{-1})$, but a precise estimate of the s^{-1} term would allow us to compute the factor for the optimal sample size

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